

No solution yet – few studies and heterogeneity

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M_HH

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Introduction

Meta-analysis is a powerful tool to summarize individual studies

Various different methods

- Frequentist approaches and Bayesian approaches
- Fixed and Random effects approaches

Often it is thought that meta-analysis is

- used to summarize **many** studies
- only **explanatory secondary** research

Meta-analyses are not only explanatory!

- Health Technology Assessment (HTA)
- Drug approval (e.g. in rare diseases)

Introduction

In general, problems arise with

- Heterogeneity between studies
- Too few studies (rare diseases!)
 - Asymptotic properties of methods
 - Probably increased heterogeneity



<http://media.propertycasualty360.com/propertycasualty360/article/2011/11/07/applesorganges1172011-crop-600x338.jpg>

Overview

Andrea Smith: Meta-Analysis with 2 studies

Theodor Framke: Heterogeneity in Meta-Analysis with few studies

Martina Kottas: Triggers of heterogeneity – alternative detection rules

Kristina Weber: Extrapolation – adult to pediatric population

Comparison of methods for MA

Simulation study to analyze frequentist meta-analysis methods for $k=2,6$

- Fixed effect (inverse variance, (FE))
- Random effects (inverse variance, DerSimonian and Laird, (DL))
- Hartung and Knapp (DL-estimator, (HK))

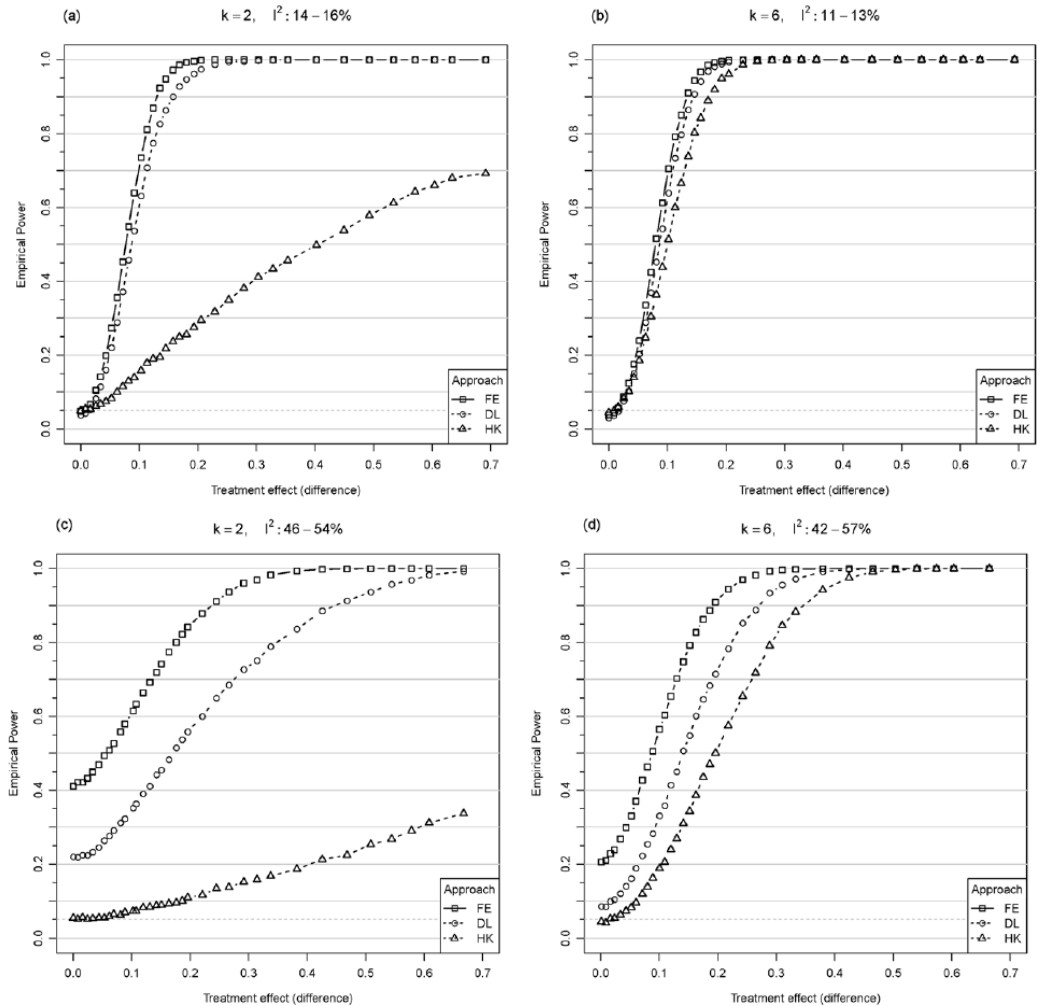
Extension

- Mantel-Haenszel method (MH), default in RevMan
- Should perform better than inverse variance method when data are sparse
- Some attention has been drawn to alternatives to estimate between-study variance for RE meta-analysis
 - E.g. Veroniki et al. (2016) lists 16 different estimators for the between-study variance
 - Paule-Mandel (PM) seems promising (e.g. Novianti et al. 2014; Langan et al. 2015; Langan et al. 2016)

Results

Alternative hypothesis for treatment effect and no heterogeneity

Alternative hypothesis for treatment effect and heterogeneity



Conclusion

Result: *There is still no solution...*

- Type I error increases when heterogeneous studies are summarized
- HK-approach is relatively safe, but lacks power for $k < 5$
- Comparable problems arise in stratified studies as methods are analogues
 → Type I error in one study with heterogeneous subgroups increases as well
 → careful assessment of subgroups required
- Summarizing only homogeneous studies with FE-approach?
- Conclusions do not change w.r.t. PM, MH

→ Detecting heterogeneity is not easy!

Statistics
in Medicine

Commentary

(wileyonlinelibrary.com) DOI: 10.1002/sim.6473

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No solution yet for combining two independent studies in the presence of heterogeneity

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Triggers/warning signals

If treatment effect is significant (based on superiority trial):

I. Q-rule: $p\text{-value of Cochran's } Q \leq 0.15$

II. Regulator's rule: $OR_i < \exp\left(\frac{\log(OR)}{2}\right)$ or $OR_i > \exp(2 \cdot \log(OR))$

III. Epidemiologist's rule: $OR_i < \exp\left(\frac{\log(OR)}{4}\right)$ or $OR_i > \exp(4 \cdot \log(OR))$

IV. CI-rule: the point estimate of a subgroup is not included in the confidence interval of the overall treatment effect

Simulation Results

- **Q-rule...**

- + Type I error as specified
- + Performance well in balanced SGs
- in unbalanced SGs power loss

- **Regulator's rule...**

- + In balanced SGs small Type I error and power similar to Q-rule
- In unbalanced SGs high Type I error

- **Epidemiologist's rule...**

- + Small type I error
- No more than 55% power

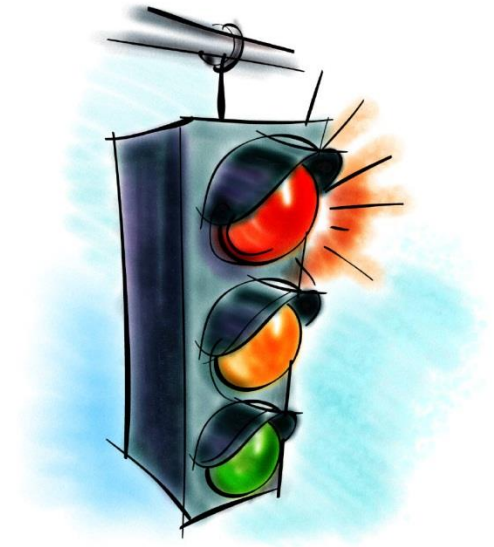
- **CI-rule...**

- + in balanced SGs small type I error
- In unbalanced SGs high type I error

Situation	SG1: SG2	Q-rule (I)	R-rule (II)	Epi-rule (III)	CI-rule (IV)
H0	50:50	0.1534	0.0686	0.0164	0.0530
H1 True II		0.5483	0.5150	0.2702	0.3693
H1 True III		0.7582	0.7419	0.4900	0.6400
H0	70:30	0.1484	0.1832	0.0608	0.2079
H1 True II		0.3449	0.5201	0.3356	0.4367
H1 True III		0.5177	0.6939	0.5143	0.6197
H0	90:10	0.1440	0.4585	0.1808	0.5308
H1 True II		0.2031	0.5952	0.4011	0.6102
H1 True III		0.3026	0.6983	0.5526	0.7058

Conclusion

- There is no perfect rule until now
- Signals should be properly understood
 (“yellow traffic light“)
- Cannot be perfect, some false positives need to be accepted



Leitthema

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Biometrische Entscheidungs- unterstützung in Zulassung und Nutzenbewertung am Beispiel der Implikationen von heterogenen Ergebnissen in Untergruppen der Studienpopulation

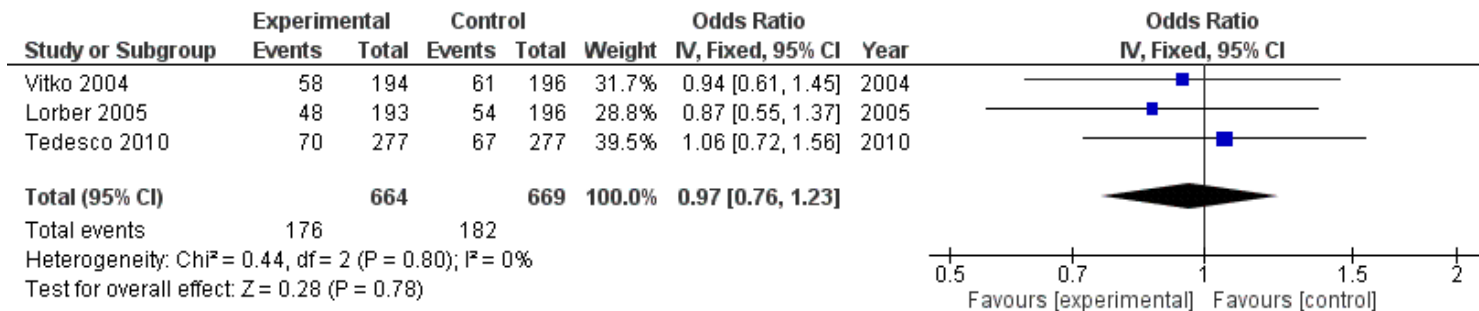
Auch um dem Vorwurf zu begegnen, dass
in einer randomisierten klinischen Studie
neue Arzneimittel in einem „Windkanal-

die möglichst größte Population zu iden-
tifizieren, für die die vorgenannten Kri-
terien erfüllt sind. Es ist jedoch quasi der

diendaten begründet wird, auffällig. Vie-
le Unterschiede in den Entscheidungsstrat-
egien bestehen dabei nur scheinbar, weil

Pediatric extrapolation

Adult studies in de novo kidney transplants with EVR NIM(log(OR)): 0.54



Aim: extrapolation to the paediatric population with one study

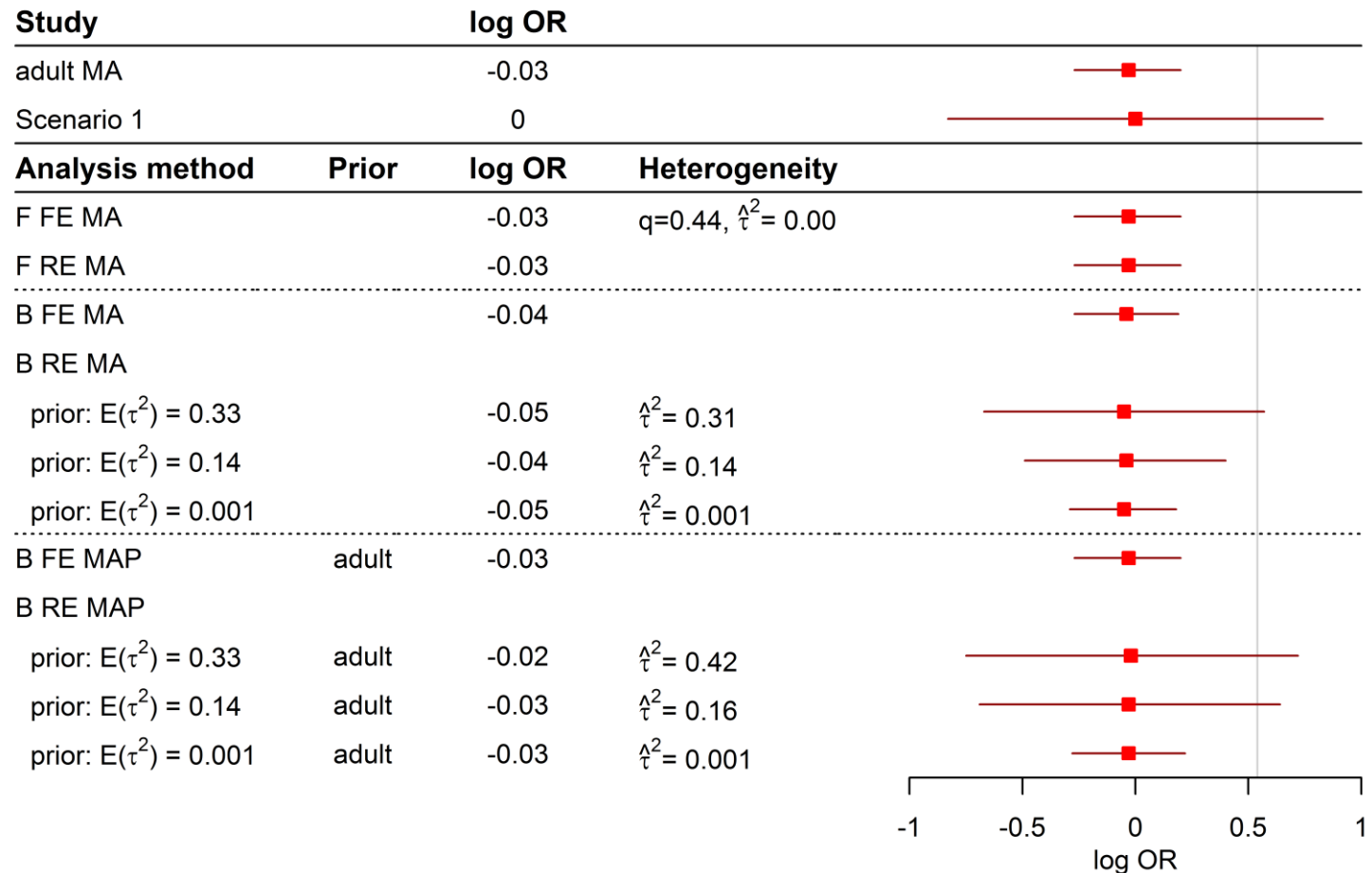
Investigation of two different scenarios:

study	EVR events/treated	MPA events/treated	log (OR) 95% CI P-value
Scenario 1	16/53 30.2%	16/53 30.2%	0.00 (-0.83; 0.83) 1.00
Scenario 2	22/53 41.5%	16/53 30.2%	0.50 (-0.31; 1.30) 0.33

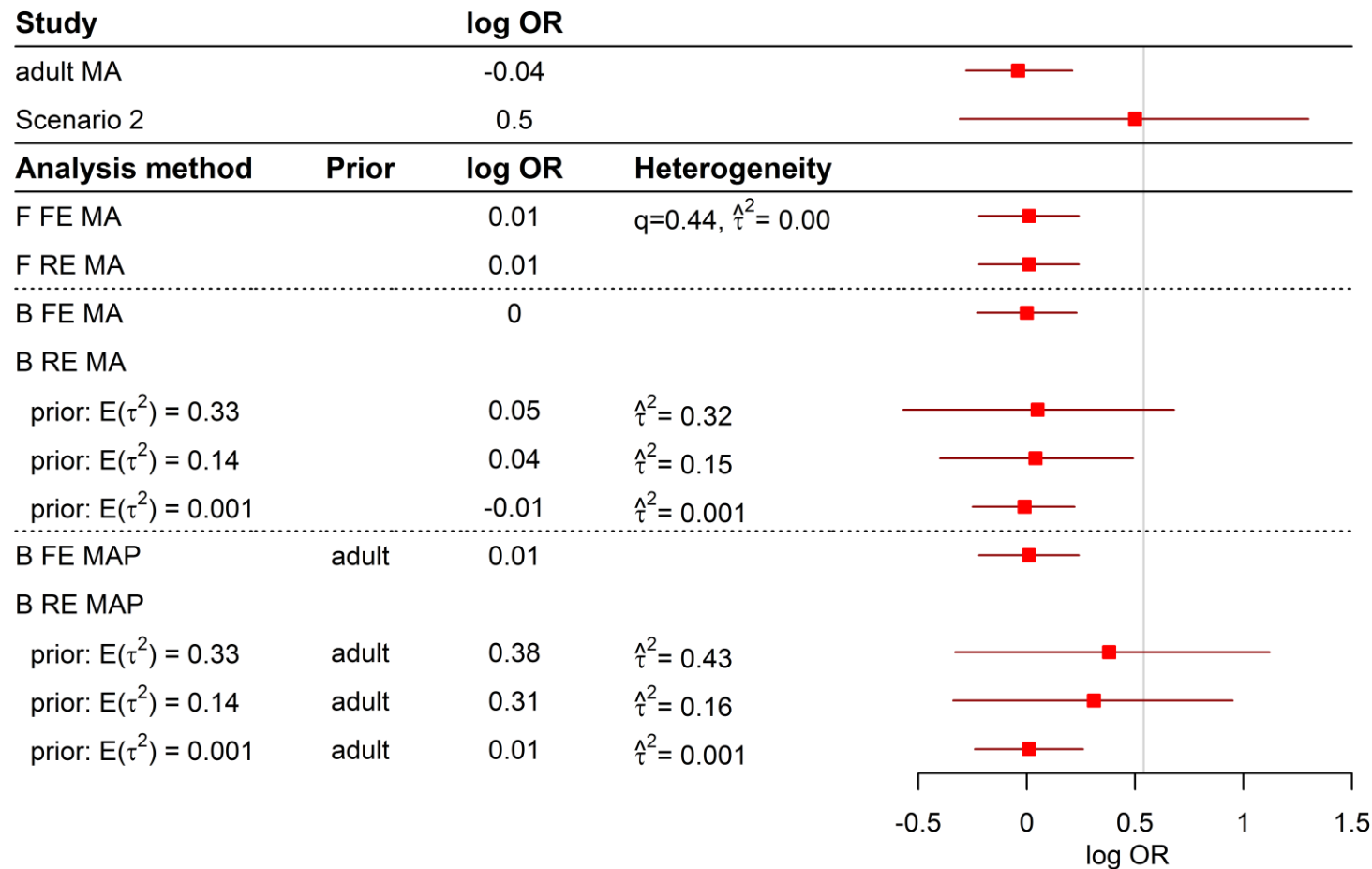
Approaches to a summary evaluation of individual sources of information

- **Frequentist Meta-Analysis**
 - Joint analysis of existing and new trial (eventually looking into heterogeneity) in a fixed (FEM) or a random (REM) effects model
- **Bayesian Meta-Analysis**
 - Joint analysis of existing and new trial in a FEM or a REM (Smith et al., 1995)
- **Bayesian meta-analytic predictive approach**
 - Analysis of new trial „in light of“ the already existing trial in a FEM or a REM (Viele et al., 2014 and Spiegelhalter et al., 2004)

Results with Scenario 1 (assumed homogeneity)



Results with Scenario 2 (log OR = 0.50, at the margin)



Assessment of the exemplary analyses

Many approaches ...

- If meta-analysis is used as a tool to arrive at an overall conclusion, no difference between a frequentist approach or a Bayesian approach can be detected: actually summary estimates will always be dominated by adult data.
- Using the predictive approach might allow that the pediatric data stand against the adult data (in case a prior is chosen that will allow for heterogeneity), however then even in case of homogeneity nothing can be concluded with the current sample-size.
- If heterogeneity is restricted, the impact of the adult data is increased (similar to frequentist MA).
- Precise pre-specification of the assumptions is required / recommended.
- Such considerations could be used to determine sample-size for a pediatric trial.

Discussion

How to **summarize**?

- Homogeneity: Fixed effect approach preferred
- Hartung-Knapp REM good T1E control. Lacks power for small k in homogeneous situation. Heterogeneity: interpretation still problematic
- Extension to MH, PM: also no solution yet
- No optimal rule available for detection of heterogeneity (false positives vs. overlooking heterogeneity?)

Extrapolation: What can be done?

- Avoiding “overweight” in the MA-approach (e.g. with content-wise selection of adult patients, only use data from young adults to weigh in for the assessment of adolescent pediatric patients)
- Be precise about the weight of the prior information
- Change of emphasis from “Does it work?” towards “Is there evidence for differential effects?”

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The holy grail of heterogeneity assessment has not been found yet – but we will continue searching.



Thank you for your attention!

Results

Null hypothesis for treatment effect, **no heterogeneity**

k	N	Sample Size Study _i : study _k	MH	FE	DL	PM	HK	Q	Mean I ²
2	480	240:240	0,0501	0,0480	0,0355	0,0355	0,0496	0,1558	15,50
2	60	30:30	0,0355	0,0241	0,0229	0,0229	0,0523	0,1373	14,41
2	480	120:360	0,0470	0,0449	0,0349	0,0349	0,0495	0,1518	15,05
6	480	80:80:....:80	0,0489	0,0372	0,0308	0,0308	0,0422	0,1313	12,09
6	180	30:30:....:30	0,0428	0,0199	0,0182	0,0181	0,0375	0,0651	8,26
6	480	60:....:60:180	0,0495	0,0358	0,0292	0,0285	0,0413	0,1167	11,33

Effect size: Odds ratio. Baseline risk: 0.2

Results

Null hypothesis for treatment effect, but **heterogeneity**

k	N	Sample Size Study _i : study _k	MH	FE	DL	PM	HK	Q	Mean I ²
2	480	240:240	0,4320	0,4026	0,2096	0,2096	0,0485	0,5618	50,00
2	60	30:30	0,6007	0,4518	0,2199	0,2199	0,0475	0,5669	50,27
2	480	120:360	0,5327	0,5206	0,2408	0,2408	0,0573	0,5637	50,27
6	480	80:80:....:80	0,2557	0,2046	0,0892	0,0860	0,0411	0,7310	50,58
6	180	30:30:....:30	0,3051	0,1993	0,0941	0,0901	0,0488	0,7355	50,72
6	480	60:....:60:180	0,3295	0,2854	0,0983	0,0932	0,0517	0,7185	50,08

Effect size: Odds ratio. Baseline risk: 0.2